Statistics II Lesson 1. Inference on one population

Year 2009/10

Lesson 1. Inference on one population

Contents

- Introduction to inference
- Point estimators
 - The estimation of the mean and variance
- Estimating the mean using confidence intervals
 - Confidence intervals for the mean of a normal population with known variance

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- Confidence intervals for the mean from large samples
- Confidence intervals for the mean of a normal population with unknown variance
- Estimating the variance using confidence intervals
 - Confidence intervals for the variance of a normal population

Lesson 1. Inference on one population

Learning goals

- Know how to estimate population values of means, variances and proportions from simple random samples
- Know how to construct confidence intervals for the mean of one population
 - In the case of a normal distribution
 - In the general case for large samples
- Know how to construct confidence intervals for the population proportion from large samples
- Know how to construct confidence intervals for the variance of one normal population

Lesson 1. Inference on one population

Bibliography references

- ▶ Meyer, P. "Probabilidad y aplicaciones estadísticas" (1992)
 - Chapter 14
- ▶ Newbold, P. "Estadística para los negocios y la economía" (1997)

Chapters 7 and 8 (up to 8.6)

Inference

Definitions

- Inference: the process of obtaining information corresponding to unknown population values from sample values
- Parameter: an unknown population value that we wish to approximate using sample values
- Statistic: a function of the information available in the sample
- Estimator: a random variable that depends on sample information and whose value approximates the value of the parameter of interest

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 Estimation: a concrete value for the estimator associated to a specific sample

Inference

Example

We wish to estimate the average yearly household food expenditure in a given region from a sample of 200 households

- The parameter of interest would be the average value of the expenditure in the region
- ► A relevant statistic in this case would be the sum of all expenditures of the households in the sample
- A reasonable estimator would be the average household expenditure in a sample
- If for a given sample the average food expenditure is 3.500 euros, the estimation of the average yearly expenditure in the region would be 3.500 euros.

- Population parameters of interest:
 - mean or variance of a population, or the proportion in the population possesing a given characteristic
- Selecting an estimator:
 - Intuitively: for example, from equivalent values in the sample
 - Or alternatively those estimators having the best properties

Properties of point estimators

- Bias: the difference between the mean of the estimator and the value of the parameter
 - If the parameter of interest is μ and the estimator is $\hat{\mu}$, its bias is defined as

 $\mathsf{Bias}[\hat{\mu}] = \mathsf{E}[\hat{\mu}] - \mu$

- Unbiased estimators: those having bias equal to zero
 - If the parameter is the population mean μ , the sample mean \bar{X} has zero bias

Properties of point estimators

- Efficiency: the value of the variance for the estimator
 - A measure related to the precision of the estimator
 - An estimator is more efficient than others if its variance is smaller
 - Relative efficiency for two estimators of a given parameter, $\hat{\theta}_1$ and $\hat{\theta}_2$,

Relative efficiency =
$$\frac{Var[\hat{\theta}_1]}{Var[\hat{\theta}_2]}$$

Comparing estimators

- Best estimator: a minimum variance unbiased estimator
 - Not always known
- Selection criterion: mean squared error
 - A combination of the two preceding criteria
 - The mean squared error (MSE) of an estimator $\hat{\theta}$ is defined as

$$\mathsf{MSE}[\hat{\theta}] = \mathsf{E}[(\hat{\theta} - \theta)^2] = \mathsf{Var}[\hat{\theta}] + (\mathsf{Bias}[\hat{\theta}])^2$$

Selecting estimators

- Minimum variance unbiased estimators
 - The sample mean for a sample of normal observations
 - The sample variance for a sample of normal observations
 - The sample proportion for a sample of binomial observations
- ▶ If an estimator with good properties is not known in advance
 - General procedures to define estimators with reasonable properties

- Maximum likelihood
- Method of moments

Exercise 1.1

From a sample of units of a given product sold in eight days,

8 6 11 9 8 10 5 7

- obtain point estimations for the following population parameters: mean, variance, standard deviation, proportion of days with sales above 7 units
- if the units sold during another six-day period have been

9 8 9 10 7 10

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compute an estimation for the difference of the means in the units sold during both periods

Results

$$\bar{x} = (8+6+11+9+8+10+5+7)/8 = 8$$

$$s^{2} = \frac{1}{n-1} \sum_{i} (x_{i} - \bar{x})^{2} = 4, \quad s = \sqrt{s^{2}} = 2$$

$$\hat{p} = (1+0+1+1+1+1+0+0)/8 = 0,625$$

$$\bar{x} - \bar{y} = 8 - (9+8+9+10+7+10)/6 = -0,833$$

Exercise 1.1

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$$\bar{x}-\bar{y} = 8-(9+8+9+10+7+10)/6 = -0,833$$

Estimation using confidence intervals

Motivation

- In many practical cases the information corresponding to a point estimate is not enough
 - It is also important to have information related to the error size
 - ▶ For example, an estimate of annual growth of 0,5 % would have very different implications if the correct value may vary between 0,3 % and 0,7 %, or if this value may be between -1,5 % and 3,5 %
 - In these cases we may wish to know some information related to the precision of the point estimator
- The most usual way to provide this information is to compute an interval estimator
 - Confidence interval: a range of values that includes the correct value of the parameter of interest with high probability

Estimation using confidence intervals

Concept

- Interval estimator
 - A rule based on sample information
 - That provides an interval containing the correct value of the parameter
 - With high probability
- For a parameter θ, given a value 1 − α between 0 and 1, the confidence level, an interval estimator is defined as two random variables θ̂_A and θ̂_B satisfying

$$P(\hat{\theta}_A \leq \theta \leq \hat{\theta}_B) = 1 - \alpha$$

- For two concrete values of these random variables, a and b, we obtain an interval [a, b] that we call a confidence interval at the 100(1 − α) % level for θ
 - 1α is known as the confidence level of the interval
 - If we generate many pairs a and b using the rule defining the interval estimator, it holds that θ ∈ [a, b] for 100(1 − α) % of the pairs (but not always)

General comments

- The confidence interval is associated to a given probability, the confidence level
- From the definition of the values defining the interval estimator, $\hat{\theta}_A$ y $\hat{\theta}_B$,

$$P(\hat{\theta}_A \leq \theta \leq \hat{\theta}_B) = 1 - \alpha$$

these values could be obtained from the values of quantiles corresponding to the distribution of the estimator, $\hat{\theta}$

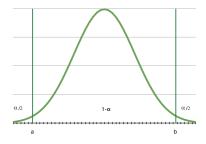
- We need to know the distribution of a quantity that relates θ and $\hat{\theta}$, in order to compute these quantiles
- This distribution is the basis for the computation of confidence intervals. It depends on
 - The parameter we wish to estimate (mean, variance)
 - The population distribution
 - ► The information that may be available (for example, if we know the value of other parameters)
- We study in this lesson different particular cases (for different parameters, distributions)

Hypotheses and goal

- We consider first a particularly simple, although not very realistic, case
- We assume that
 - we have a simple random sample of n observations
 - the population follows a normal distribution
 - \blacktriangleright we know the population variance σ^2
- \blacktriangleright Goal: construct a confidence interval for the (unknown) population mean μ
 - For a confidence level 1α , either prespecified or selected by us

The mean of a normal population with known variance Procedure

- ▶ Let X₁,..., X_n denote the simple random sample and X̄ its sample mean, our point estimator
- \blacktriangleright Our first step is to obtain information on the distribution of a variable that relates μ and \bar{X}
- From this distribution we obtain a pair of values a and b that define the interval (for the variable) having the desired probability
- \blacktriangleright From these values we define an interval for μ



Procedure

For the case under consideration, the distribution of the sample mean satisfies _

$$rac{ar{X}-\mu}{\sigma/\sqrt{n}}\sim N(0,1)$$

- A known distribution
- Relating \bar{X} and μ
- We construct an interval containing the desired probability for a standard normal distribution, finding a value z_{α/2} satisfying

$$P(-z_{\alpha/2} \leq Z \leq z_{\alpha/2}) = 1 - \alpha$$

• $z_{\alpha/2}$ is the value such that a standard normal distribution takes larger values with probability equal to $\alpha/2$

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Procedure

The following interval has the desired probability

$$-z_{\alpha/2} \leq \frac{\bar{X}-\mu}{\sigma/\sqrt{n}} \leq z_{\alpha/2}$$

 \blacktriangleright Replacing the sample values and solving for μ we obtain the desired confidence interval

$$\bar{x} - z_{\alpha/2} \frac{\sigma}{\sqrt{n}} \le \mu \le \bar{x} + z_{\alpha/2} \frac{\sigma}{\sqrt{n}}$$

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Exercise 1.2

A bottling process for a given liquid produces bottles whose weight follows a normal distribution with standard deviation equal to 55 gr. A simple random sample of 50 bottles has been selected; its mean weight has been 980 gr. Compute a confidence interval at 99% for the mean weight of all bottles from the process

Results

$$\begin{array}{rcl} \displaystyle \frac{\bar{X}-\mu}{55/\sqrt{50}} & \sim & \textit{N}(0,1) \\ \\ \alpha & = & 1-0,99 = 0,01, \qquad z_{\alpha/2} = z_{0,005} = 2,576 \\ \\ & -2,576 \leq \frac{980-\mu}{55/\sqrt{50}} \leq 2,576 \\ \\ & 959,96 \leq \mu \leq 1000,04 \end{array}$$

Exercise 1.2

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General procedure

Steps:

- 1. Identify the variable with a known distribution and the distribution we will use to construct the confidence interval
- 2. Find the percentiles of the distribution that correspond to the selected confidence level
- 3. Construct the interval for the variable with known distribution
- 4. Replace the sample values in the interval
- 5. Solve for the value of the parameter in the interval, to obtain another interval, specific for this parameter

Properties of the interval

- The size of the confidence interval is a measure of the precision in the estimation
- In the preceding case this size was given as

$$\frac{2z_{\alpha/2}\sigma}{\sqrt{n}}$$

Thus, the precision depended on

- ▶ The standard deviation of the population. The larger it is, the lower the precision in the estimation
- The sample size. The precision increases as the size increases
- The confidence level. If we select a higher level we obtain a larger interval

Exercise 1.3

For the data in exercise 1.2, compute the changes in the confidence interval if

- the sample size increases to 100 (for the same value of the sample mean)
- \blacktriangleright the confidence level is modified to 95 %

Results

$$\begin{aligned} -2,576 &\leq \frac{980-\mu}{55/\sqrt{100}} \leq 2,576 \\ 965,83 &\leq \mu \leq 994,17 \\ \alpha &= 1-0,95 = 0,05, \qquad z_{\alpha/2} = z_{0,025} = 1,96 \\ -1,96 &\leq \frac{980-\mu}{55/\sqrt{50}} \leq 1,96 \\ 964,75 &\leq \mu \leq 995,25 \end{aligned}$$

Exercise 1.3

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Motivation

- In many practical cases we do not know if the population distribution is normal or the value of its standard deviation
- If the sample size is sufficiently large, the central limit theorem allows us to construct approximate confidence intervals

Hipotheses and goal

- We assume that
 - we have a simple random sample of size n
 - the size of the sample is sufficiently large
- Goal: construct an approximate confidence interval for the (unknown) population mean μ
 - For a confidence level 1α , either prespecified or selected by us

Procedure

▶ For the case we are considering, the central limit theorem specifies that for *n* large enough it holds approximately that

$$rac{ar{X}-\mu}{S/\sqrt{n}}\sim N(0,1)$$

where S denotes the sample standard deviation

- It is the same distribution as in the preceding case
- It provides a relationship between \bar{X} y μ
- ► We build as before an interval that contains the desired probability under a standard normal distribution, for a value $z_{\alpha/2}$ satisfying

$$P(-z_{\alpha/2} \leq Z \leq z_{\alpha/2}) = 1 - \alpha$$

Procedure

The following interval has the desired probability under the distribution _____

$$-z_{\alpha/2} \leq \frac{\bar{X}-\mu}{S/\sqrt{n}} \leq z_{\alpha/2}$$

 \blacktriangleright Replacing the sample values and solving for the value of μ in the inequalities we obtain the desired confidence interval

$$\bar{x} - z_{\alpha/2} \frac{s}{\sqrt{n}} \le \mu \le \bar{x} + z_{\alpha/2} \frac{s}{\sqrt{n}}$$

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Exercise 1.4

A survey has been conducted on 60 persons. Each person provided a rating between 0 and 5 corresponding to their perception of the quality of a given service. The average rating from the sample was 2.8 and the sample standard deviation was 0.7. Compute a confidence interval at the 90% level for the average rating of the service in the population

Results

$$lpha = 1 - 0.9 = 0.1, \qquad z_{lpha/2} = z_{0.05} = 1,645$$

 $-1.645 \le \frac{2.8 - \mu}{0.7/\sqrt{60}} \le 1.645$
 $2.65 \le \mu \le 2.95$

Exercise 1.4

A survey has been conducted on 60 persons. Each person provided a rating between 0 and 5 corresponding to their perception of the quality of a given service. The average rating from the sample was 2.8 and the sample standard deviation was 0.7. Compute a confidence interval at the 90% level for the average rating of the service in the population

Results

$$egin{array}{rcl} lpha &=& 1-0,9=0,1, & z_{lpha/2}=z_{0,05}=1,645 \ & -1,645\leq rac{2,8-\mu}{0,7/\sqrt{60}}\leq 1,645 \ & 2,65\leq \mu\leq 2,95 \end{array}$$

Motivation

- We wish to estimate the proportion of a population that satisfies a certain condition, from sample data
- Estimating proportions is a particular case of the preceding one when we had nonnormal data
- Our estimator in this case will be the sample proportion
- ▶ If X_i represents whether a member of the simple random sample of size n satisfies the condition, or does not satisfy it, and the probability of satisfaction is p, then X_i follows a Bernoulli distribution
- ▶ We wish to estimate *p*, the proportion in the population that satisfies the condition

- Using the sample proportion $\hat{p} = \sum_i X_i / n$
 - *p̂* is a sample mean

Hypotheses and goal

- We assume that
 - ▶ we have a simple random sample of size *n*, where each observation takes either the value 0 or 1
 - the sample size is sufficiently large
- Goal: construct an approximate confidence interval for the (unknown) population proportion p
 - For a confidence level 1α , either prespecified or selected by us

Procedure

- ▶ In our case, compared with the preceding one, $\mu = p$, $\sigma^2 = p(1-p)$
- The central limit theorem states that for large n it holds approximately that

$$rac{\hat{p}-p}{\sqrt{p(1-p)/n}}\sim N(0,1)$$

where \hat{p} denotes the proportion in the sample

- We approximate p(1-p) with the corresponding sample value $\hat{p}(1-\hat{p})$
 - The resulting variable follows approximately the same distribution
- We construct, as before, an interval containing the desired probability for a standard normal distribution, computing a value z_{α/2} satisfying

$$P(-z_{\alpha/2} \leq Z \leq z_{\alpha/2}) = 1 - \alpha$$

Procedure

▶ The following interval has the desired probability

$$-z_{\alpha/2} \leq rac{\hat{p}-p}{\sqrt{\hat{p}(1-\hat{p})/n}} \leq z_{\alpha/2}$$

Replacing sample values and solving for the value of p in the inequalities we obtain the desired confidence interval

$$\hat{p} - z_{\alpha/2}\sqrt{rac{\hat{p}(1-\hat{p})}{n}} \leq p \leq \hat{p} + z_{\alpha/2}\sqrt{rac{\hat{p}(1-\hat{p})}{n}}$$

Exercise 1.5

In a sample of 200 patients it has been observed that the number having serious complications associated to a given illness is 38. Compute a confidence interval at the 99 % level for the proportion of patients in the population that may have serious complications associated to that illness

Results

$$\begin{array}{rcl} \hat{p} & = & 38/200 = 0,19 \\ \alpha & = & 1-0,99 = 0,01, \qquad z_{\alpha/2} = z_{0,005} = 2,570 \\ & -2,576 \leq \frac{0,19-p}{\sqrt{0,19(1-0,19)/200}} \leq 2,576 \\ & 0,119 \leq \mu \leq 0,261 \end{array}$$

Exercise 1.5

In a sample of 200 patients it has been observed that the number having serious complications associated to a given illness is 38. Compute a confidence interval at the 99 % level for the proportion of patients in the population that may have serious complications associated to that illness

Results

$$\begin{array}{rcl} \hat{p} & = & 38/200 = 0,19 \\ \alpha & = & 1-0,99 = 0,01, \qquad z_{\alpha/2} = z_{0,005} = 2,576 \\ & & -2,576 \leq \frac{0,19-p}{\sqrt{0,19(1-0,19)/200}} \leq 2,576 \\ & & 0,119 \leq \mu \leq 0,261 \end{array}$$

Motivation

- We wish to estimate the mean of the population
- And we know that the distribution in the population is normal
- But we do not know the variance of the population
- ▶ If the sample size is small the preceding results would not be applicable
- But for this particular case we know the distribution of the sample mean for any sample size

Hypotheses and goal

- We assume that
 - we have a simple random sample of size n
 - the population follows a normal distribution
- \blacktriangleright Goal: construct a confidence interval for the (unknown) population mean μ
 - For a confidence level 1α , either prespecified or selected by us

Procedure

In this case the basic distribution result is

$$rac{ar{X}-\mu}{S/\sqrt{n}}\sim t_{n-1}$$

where S denotes the sample standard deviation and t_{n-1} denotes the Student-t distribution with n-1 degrees of freedom

- A symmetric distribution (around zero) similar to the normal one (it converges to a normal distribution as *n* increases)
- ► We again construct an interval containing the desired probability, but we do it for a Student-*t* distribution with *n* − 1 degrees of freedom, finding a value *t*_{*n*−1,α/2} satisfying

$$P(-t_{n-1,\alpha/2} \leq T_{n-1} \leq t_{n-1,\alpha/2}) = 1 - \alpha$$

Procedure

The following interval corresponds to the required probability

$$-t_{n-1,\alpha/2} \leq \frac{\bar{X}-\mu}{S/\sqrt{n}} \leq t_{n-1,\alpha/2}$$

Replacing sample values and solving for the value of µ in the inequalities we obtain the desired confidence interval

$$\bar{x} - t_{n-1,\alpha/2} \frac{s}{\sqrt{n}} \le \mu < \bar{x} + t_{n-1,\alpha/2} \frac{s}{\sqrt{n}}$$

Exercise 1.6

You have measured the working life of a sample of 20 high-efficiency light bulbs. For this sample the measured mean life has been equal to 4520 h., with a sample standard deviation equal to 750 h. If the working life of these bulbs is assumed to follow a normal distribution, compute a confidence interval at the 95 % level for the average life of all the bulbs (population mean)

Results

 $\begin{array}{ll} \alpha & = & 1-0,95 = 0,05, & t_{n-1,\alpha/2} = t_{19,0,025} = 2,093 \\ & -2,093 \leq \frac{4520-\mu}{750/\sqrt{20}} \leq 2,093 \\ & 4169,0 \leq \mu \leq 4871,0 \end{array}$

Exercise 1.6

You have measured the working life of a sample of 20 high-efficiency light bulbs. For this sample the measured mean life has been equal to 4520 h., with a sample standard deviation equal to 750 h. If the working life of these bulbs is assumed to follow a normal distribution, compute a confidence interval at the 95 % level for the average life of all the bulbs (population mean)

Results

$$egin{array}{rcl} lpha &=& 1-0,95=0,05, & t_{n-1,lpha/2}=t_{19,0,025}=2,093 \ && -2,093 \leq rac{4520-\mu}{750/\sqrt{20}} \leq 2,093 \ && 4169,0 \leq \mu \leq 4871,0 \end{array}$$

Motivation

- Up to now we have only considered confidence intervals for the population mean
- In some cases we may also be interested in knowing confidence intervals for the variance
- ▶ The relevant distributions are not known in general, except for a few cases
- We will only consider the case of normal data

Hypotheses and goal

- We assume that
 - we have a simple random sample of size n
 - the population follows a normal distribution
- \blacktriangleright Goal: Goal: construct a confidence interval for the (unknown) population variance σ^2
 - For a confidence level 1α , either prespecified or selected by us

Procedure

In this case the basic result is

$$\frac{(n-1)S^2}{\sigma^2} \sim \chi^2_{n-1}$$

where S denotes the sample standard deviation and χ^2_{n-1} denotes the chi-squared distribution with n-1 degrees of freedom

- It is an asymmetric distribution that takes nonnegative values
- As in the preceding cases, the first step is to construct an interval that has the desired probability under the chi-squared distribution
 - ► As the chi-squared distribution is asymmetric, we need two values to define the interval, $\chi^2_{n-1,1-\alpha/2}$ and $\chi^2_{n-1,\alpha/2}$

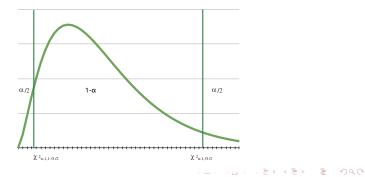
Variance of a normal population Procedure

We select these values from the conditions

$$P(\chi^2_{n-1} \ge \chi^2_{n-1,1-\alpha/2}) = 1 - \alpha/2, \qquad P(\chi^2_{n-1} \ge \chi^2_{n-1,\alpha/2}) = \alpha/2$$

They satisfy

$$P(\chi^{2}_{n-1,1-\alpha/2} \le \chi^{2}_{n-1} \le \chi^{2}_{n-1,\alpha/2}) = 1 - \alpha$$



Procedure

▶ The following interval corresponds to the desired probability

$$\chi^2_{n-1,1-\alpha/2} \le \frac{(n-1)S^2}{\sigma^2} \le \chi^2_{n-1,\alpha/2}$$

Replacing sample values and solving for σ² in the inequalities we obtain the interval

$$\frac{(n-1)s^2}{\chi^2_{n-1,\alpha/2}} \leq \sigma^2 \leq \frac{(n-1)s^2}{\chi^2_{n-1,1-\alpha/2}}$$

 For the standard deviation the corresponding confidence interval will be given by

$$\sqrt{\frac{(n-1)s^2}{\chi^2_{n-1,\alpha/2}}} \le \sigma \le \sqrt{\frac{(n-1)s^2}{\chi^2_{n-1,1-\alpha/2}}}$$

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Exercise 1.7

For the data in exercise 1.6, you are asked to compute a confidence interval at the 95 % level for the (population) standard deviation of the bulb life

Results

$$\alpha = 1 - 0.95 = 0.05$$

$$\chi^{2}_{n-1,1-\alpha/2} = \chi^{2}_{19,0,975} = 8,907$$

$$\chi^{2}_{n-1,\alpha/2} = \chi^{2}_{19,0,025} = 32,852$$

$$8,907 \le \frac{19 \times 750^{2}}{\sigma^{2}} \le 32,852$$

$$325323 \le \sigma^{2} \le 1199899$$

$$570.37 \le \sigma \le 1095.40$$

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Exercise 1.7

For the data in exercise 1.6, you are asked to compute a confidence interval at the 95 % level for the (population) standard deviation of the bulb life

Results

$$\begin{array}{rcl} \alpha &=& 1-0.95 = 0.05 \\ \chi^2_{n-1,1-\alpha/2} &=& \chi^2_{19,0,975} = 8.907 \\ \chi^2_{n-1,\alpha/2} &=& \chi^2_{19,0,025} = 32.852 \\ && 8.907 \leq \frac{19 \times 750^2}{\sigma^2} \leq 32.852 \\ && 325323 \leq \sigma^2 \leq 1199899 \\ && 570.37 \leq \sigma \leq 1095.40 \end{array}$$

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Confidence intervals

Summary for one population

► For a simple random sample

Parameter	Hipotheses	Distribution	Interval
	Normal data Known variance	$\frac{\bar{X}-\mu}{\sigma/\sqrt{n}} \sim N(0,1)$	$\mu \in \left[\bar{x} - z_{\alpha/2} \frac{\sigma}{\sqrt{n}}, \bar{x} + z_{\alpha/2} \frac{\sigma}{\sqrt{n}}\right]$
Mean	Nonnormal data Large sample	$rac{ar{X}-\mu}{S/\sqrt{n}}\sim N(0,1)$	$\mu \in \left[\bar{x} - z_{\alpha/2} \frac{s}{\sqrt{n}}, \bar{x} + z_{\alpha/2} \frac{s}{\sqrt{n}}\right]$
	Proportions Large sample	$\frac{\hat{P}-p}{\sqrt{\hat{P}(1-\hat{P})/n}} \sim N(0, 1)$	$p \in \left[\hat{p} - z_{\alpha/2} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}, \hat{p} + z_{\alpha/2} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}\right]$
	Normal data Unknown var.	$\frac{\tilde{X}-\mu}{S/\sqrt{n}} \sim t_{n-1}$	$\mu \in \left[\bar{x} - t_{n-1,\alpha/2} \frac{s}{\sqrt{n}}, \bar{x} + t_{n-1,\alpha/2} \frac{s}{\sqrt{n}}\right]$
Variance	Normal data	$\frac{(n-1)S^2}{\sigma^2} \sim \chi^2_{n-1}$	$\sigma^{2} \in \left[\frac{(n-1)s^{2}}{\chi^{2}_{n-1,\alpha/2}}, \frac{(n-1)s^{2}}{\chi^{2}_{n-1,1-\alpha/2}}\right]$

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